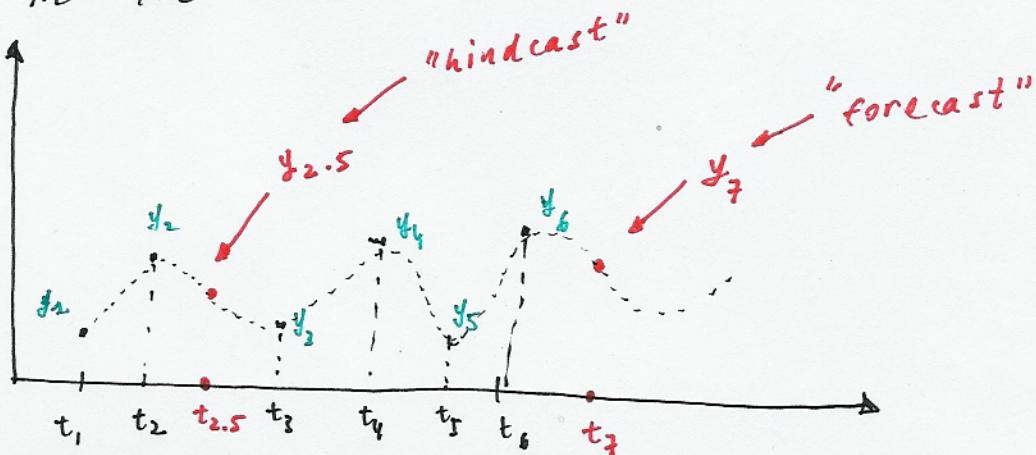


Covariance Modeling

The main idea is to use covariability between signals to interpolate, extrapolate and forecast a signal (s) of interest.

Example in the TIME DOMAIN



(Fig. 2)

Suppose you want to find the value of $y(t=t_{2.5})$.

You can use the covariance $\langle y(t) y(t-\tau) \rangle$ to interpolate $y(t=t_{2.5})$. This operation is called "hindcast" = "time interpolation". One could use this covariance model to find the value of $y(t=t_7)$ at a future time. This operation is called "forecast" = "time extrapolation".

{interpolation}: estimate the value of $y(t_n)$ within
hindcast } inside the data domain.

{extrapolation}: estimate the value of $y(t_n)$ outside
forecast } the data domain.

"Univariate formulation"

$$y = \underset{\text{predictand}}{\alpha} \underset{\text{predictor}}{x} + \underset{\text{error}}{\eta}$$
$$\langle y \rangle = \phi \quad \langle \alpha \rangle = \phi \quad \langle \eta \rangle = \phi$$
$$\alpha = \frac{\langle yx \rangle}{\langle xx \rangle} \quad \text{covariance Model also known as best "optimal linear estimator".}$$

As you know this is also the least square estimate obtained by minimizing $J = \sum_{i=1}^N \eta_i^2$

Once you predict $\hat{y} = \alpha x$, one can also predict the error $\varepsilon = \sqrt{\langle \eta^2 \rangle} \rightarrow \hat{y} \pm \varepsilon$

where

$$\langle \eta^2 \rangle = \langle y^2 \rangle (1 - e^2)$$

$$e^2 = \frac{\langle xy \rangle^2}{\langle x^2 \rangle \langle y^2 \rangle}$$

correlation coefficient
square = fraction of variance that is explained by the model.

The covariance modeling is a very general statement that reappears over and over in linear statistics with a wide range of applications and with different flavours.

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"Multivariate Formulation"

\underline{y} = vector of predictand

\underline{x} = vector of predictor

$$\underline{y} = \underline{\alpha} \underline{x} + \underline{\eta}$$

Similarly to the univariate case

$$\underline{y} \underline{x}^T = \underline{\alpha} \underline{x} \underline{x}^T + \underline{\eta} \underline{x}^T$$

$$\langle \underline{y} \underline{x}^T \rangle = \underline{\alpha} \langle \underline{x} \underline{x}^T \rangle$$

note that $(\langle \underline{y} \underline{x}^T \rangle)^T = \langle \underline{x} \underline{y}^T \rangle$

$$\underline{\alpha} = \langle \underline{y} \underline{x}^T \rangle \langle \underline{x} \underline{x}^T \rangle^{-1}$$

↑
the covariance model

Once you predict $\hat{\underline{y}} = \underline{\alpha} \underline{x}$ one can estimate
the error $\underline{\epsilon} = \sqrt{\text{diag}(\underline{\eta} \underline{\eta}^T)}$

$$\underline{\eta} \underline{\eta}^T = \begin{bmatrix} \langle \eta_1 \eta_1 \rangle & \langle \eta_1 \eta_2 \rangle & & \\ \langle \eta_2 \eta_1 \rangle & \ddots & & \\ & \ddots & \ddots & \\ & & & \langle \eta_N \eta_N \rangle \end{bmatrix}$$

diagonal elements are the error variance for each y_i

$$\langle \underline{\eta} \underline{\eta}^T \rangle = \langle (\underline{y} - \hat{\underline{y}}) (\underline{y} - \hat{\underline{y}})^T \rangle$$

$$= \langle \underline{y} \underline{y}^T \rangle + \langle \hat{\underline{y}} \hat{\underline{y}}^T \rangle - \langle \hat{\underline{y}} \underline{y}^T \rangle - \langle \underline{y} \hat{\underline{y}}^T \rangle$$

$$\langle \underline{n} \underline{n}^T \rangle = \langle \underline{y} \underline{y}^T \rangle + \cancel{\langle \underline{x} \underline{x}^T \rangle} - \cancel{\langle \underline{x} \underline{y}^T \rangle} - \cancel{\langle \underline{y} \underline{x}^T \rangle}$$

using $\cancel{\langle \cdot \cdot \cdot \rangle} = \langle \underline{y} \underline{x}^T \rangle \langle \underline{x} \underline{x}^T \rangle^{-1}$

$$\begin{aligned} \langle \underline{n} \underline{n}^T \rangle &= \langle \underline{y} \underline{y}^T \rangle + \cancel{\langle \underline{y} \underline{x}^T \rangle} \cancel{\langle \underline{x} \underline{x}^T \rangle^{-1}} \cancel{\langle \underline{x} \underline{x}^T \rangle} \underbrace{[\cancel{\langle \underline{y} \underline{x}^T \rangle} \cancel{\langle \underline{x} \underline{x}^T \rangle^{-1}}]}_{\cancel{\underline{\alpha}^T}} \\ &\quad - \cancel{\langle \underline{y} \underline{x}^T \rangle} \cancel{\langle \underline{x} \underline{x}^T \rangle^{-1}} \cancel{\langle \underline{x} \underline{y}^T \rangle} - \cancel{\langle \underline{y} \underline{x}^T \rangle} \underbrace{[\cancel{\langle \underline{y} \underline{x}^T \rangle} \cancel{\langle \underline{x} \underline{x}^T \rangle^{-1}}]}_{\cancel{\underline{\alpha}^T}} \end{aligned}$$

$$\begin{aligned} \langle \underline{n} \underline{n}^T \rangle &= \langle \underline{y} \underline{y}^T \rangle + \cancel{\langle \underline{y} \underline{x}^T \rangle} \cancel{\langle \underline{x} \underline{x}^T \rangle^{-1}} \cancel{\langle \underline{x} \underline{y}^T \rangle} \\ &\quad \text{cancel} \\ &\quad - \cancel{\langle \underline{y} \underline{x}^T \rangle} \cancel{\langle \underline{x} \underline{x}^T \rangle^{-1}} \cancel{\langle \underline{x} \underline{y}^T \rangle} - \cancel{\langle \underline{y} \underline{x}^T \rangle} \cancel{\langle \underline{x} \underline{x}^T \rangle^{-1}} \cancel{\langle \underline{x} \underline{y}^T \rangle} \end{aligned}$$

where $\cancel{\langle \underline{y} \underline{x}^T \rangle^T} = \cancel{\langle \underline{x} \underline{y}^T \rangle}$

$\cancel{[\langle \underline{x} \underline{x}^T \rangle^{-1}]}^T = \cancel{\langle \underline{x} \underline{x}^T \rangle^{-1}}$

$$\langle \underline{n} \underline{n}^T \rangle = \langle \underline{y} \underline{y}^T \rangle - \cancel{\langle \underline{y} \underline{x}^T \rangle} \cancel{\langle \underline{x} \underline{x}^T \rangle^{-1}} \cancel{\langle \underline{x} \underline{y}^T \rangle}$$

$$= \langle \underline{y} \underline{y}^T \rangle \left[1 - \cancel{\langle \underline{y} \underline{y}^T \rangle^{-1} \langle \underline{y} \underline{x}^T \rangle \cancel{\langle \underline{x} \underline{x}^T \rangle^{-1}} \cancel{\langle \underline{x} \underline{y}^T \rangle}} \right]$$

equivalent to R^2 , this is the fraction of variance explained by model.

if we define

$$\begin{aligned}\underline{\underline{C}}_{yx} &= \langle \underline{\underline{y}} \underline{\underline{x}}^T \rangle \\ \underline{\underline{C}}_{yy} &= \langle \underline{\underline{y}} \underline{\underline{y}}^T \rangle \\ \underline{\underline{C}}_{xx} &= \langle \underline{\underline{x}} \underline{\underline{x}}^T \rangle\end{aligned}$$

(1)

$$\hat{\underline{\underline{y}}} = \underline{\underline{C}}_{yx} \underline{\underline{C}}_{xx}^{-1} \underline{\underline{x}}$$

(2)

$$\langle \underline{\underline{n}} \underline{\underline{n}}^T \rangle = \underline{\underline{C}}_{yy} \left[\mathbf{1} - \underline{\underline{C}}_{yy}^{-1} \underline{\underline{C}}_{yx} \underline{\underline{C}}_{xx}^{-1} \underline{\underline{C}}_{yx}^T \right]$$

$$\langle \underline{\underline{n}} \underline{\underline{n}}^T \rangle = \underline{\underline{C}}_{yy} \left[\mathbf{1} - \left(\underline{\underline{C}}_{yy}^{-1} \underline{\underline{C}}_{yx} \right) \left(\underline{\underline{C}}_{yx} \underline{\underline{C}}_{xx}^{-1} \right)^T \right]$$

(3)

3 - steps

- ① compute covariances or prescribe a functional form of the covariance (e.g. gaussian)

- ② Make your prediction of $\hat{\underline{\underline{y}}}$ using the covariance model

$$\underline{\underline{C}}_{yx} \underline{\underline{C}}_{xx}^{-1} = \underline{\underline{\alpha}}$$

(note that if $\langle \underline{\underline{x}} \underline{\underline{x}}^T \rangle = \mathbf{I}$ then $\underline{\underline{\alpha}} \equiv \underline{\underline{C}}_{yx}$)

- ③ Estimate your error.

(note that if $\langle \underline{\underline{x}} \underline{\underline{x}}^T \rangle = \underline{\underline{C}}_{yy} = \mathbf{I}$)

$$\langle \underline{\underline{n}} \underline{\underline{n}}^T \rangle = \left[\mathbf{1} - \underline{\underline{C}}_{yx} \underline{\underline{C}}_{yx}^T \right]$$

The model

$$\hat{y} = \underbrace{\underline{\underline{C}}_{yx} \underline{\underline{C}}_{xx}^{-1}}_{\propto} x$$

WHEN TO USE:
if you know the
stats. between
 x and y

Optimal Linear Estimator
 \propto

Recall the LSP model

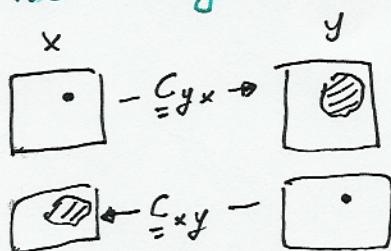
$$\hat{y} = \underline{\underline{E}} x$$

WHEN TO USE:
if you do not know
the stats. between x and y
you can assume a model E

this was the case where $\underline{\underline{E}}$ was chosen, x was known and y was unknown.

the analogy between $\underline{\underline{E}}$ and $\underline{\underline{C}}_{yx} \underline{\underline{C}}_{xx}^{-1}$ is revealing. on the except when we pick a model $\underline{\underline{E}}$ to express the relationship between y and x we are really making an assumption on the covariance matrix.

"The Adjoint" $\rightarrow \underline{\underline{E}}^T = [\underline{\underline{C}}_{yx} \underline{\underline{C}}_{xx}^{-1}]^T = \underline{\underline{C}}_{xx}^{-1} \underline{\underline{C}}_{xy}$



the adjoint is an estimate of the covariance between x and y